## Name:

1) Time to respond for 911 calls was measured for 75 calls in two cities, $C 1$ and $C 2$. The results of the study were:

| Measure | C1 | C2 |
| :--- | :---: | :---: |
| Sample size | 75 | 75 |
| Sample mean | 5.4 | 6.1 |
| Sample variance | 1.57 | 2.54 |

This problem is solved exactly like Example 8.3 in the text. a) Estimate the difference in the mean response time for the two cities.
The point estimate of $\mu_{1}-\mu_{2}$ is

$$
\bar{y}_{1}-\bar{y}_{2}=5.4-6.1=-0.7
$$

b) Find a bound for the error of estimation. See Example 8.3 in the text. The standard deviation of the estimated difference is:

$$
\sigma_{\bar{Y}_{1}-\bar{Y}_{2}}=\sqrt{\frac{\sigma_{1}^{2}}{n_{1}}+\frac{\sigma_{2}^{2}}{n_{2}}} \approx \sqrt{\frac{s_{1}^{2}}{n_{1}}+\frac{s_{2}^{2}}{n_{2}}}=\sqrt{\frac{1.57}{75}+\frac{2.54}{75}}=0.234
$$

The probability that the error of estimation is less than $\sigma_{\bar{Y}_{1}-\bar{Y}_{2}} \approx$ $2(0.234)=0.468$ is .95 . (two standard deviations)
2) An exit poll of 1,000 voters finds that 530 supported a certain candidate.
a) Estimate the proportion of the voting population that supports the candidate.

The estimator is

$$
\hat{p}=\frac{y}{n}=\frac{530}{1000}=0.530
$$

b) Find a bound for the error of estimation. The standard error for the estimator $\hat{p}$ is given by:

$$
\sigma_{\hat{p}}=\sqrt{\frac{p q}{n}} \approx \sqrt{\frac{\hat{p}(1-\hat{p})}{n}}=\sqrt{\frac{0.53 \cdot 0.47}{1000}}=0.0158
$$

With probability .95, the error of estimation is less than $2 \sigma_{\hat{p}}=0.0316$
3) Suppose $Y$ is a single observation from an exponential distribution with unknown mean $\theta$.
a) Use the method of moment-generating functions to show that $2 Y / \theta$ is a pivotal quantity having a $\chi^{2}$ distribution with two degrees of freedom.
From the table in the back of the text, the moment-generating function of $y$ is

$$
m_{y}(t)=(1-\theta t)^{-1}
$$

so the moment-generating function of $u=2 Y / \theta$ is

$$
m_{u}(t)=m_{y}\left(\frac{2}{\theta} t\right)=(1-2 t)^{-1}
$$

Comparing this to the table in the back cover of the text, we see this is the moment-generating function of a Chi-square random variable with 2 degrees of freedom.
b) Use the pivotal quantity from part a) to derive a $90 \%$ confidence interval for $\theta$. Compare your result with Example 8.4 in the text.
We need to find bounds $L$ and $U$ so that

$$
P\left(L \leq \chi_{2}^{2} \leq U\right)=0.90
$$

There are many intervals that satisfy this inequality, but only one is symmetric in the sense that the areas under the graph of the density function from zero to $L$ and from $U$ to infinity are both 0.05 .

We can find the value of $L$ and $U$ from the table in the back of the text (use the row with $\mathrm{df}=2$ ). The value of $L$ is under the column heading $\chi_{0.950}^{2}$ and is 0.102587 . The value of $U$ is under the column heading $\chi_{0.050}^{2}$ and is 5.99147. You can also obtain these from a spreadsheet with the formulas

$$
L=\operatorname{CHIINV}(0.95,2) \quad \text { and } \quad U=\operatorname{CHIINV}(0.05,2)
$$

so, for a Chi-square variable with 2 degrees of freedom,

$$
P\left(0.102587 \leq \chi_{2}^{2} \leq 5.99147\right)=0.90
$$

substituting $2 Y / \theta$ into this inequality for $\chi_{2}^{2}$ we get:

$$
P\left(0.102587 \leq \frac{2 Y}{\theta} \leq 5.99147\right)=0.90
$$

which rearranges to

$$
P\left(\frac{2 Y}{5.99147} \leq \theta \leq \frac{2 Y}{0.102587}\right)=0.90
$$

Note that this produces the same confidence interval as the one obtained in Example 8.4.
4) Now suppose $\left(Y_{1}, \ldots, Y_{9}\right)$ is a sample of size $n=9$ from an exponential distribution with mean $\theta$.
a) Use the method of moment-generating functions to show that

$$
\frac{2}{\theta} \sum_{i=1}^{9} Y_{i}
$$

is a pivotal quantity having a $\chi^{2}$ distribution with 18 degrees of freedom.

Because the $Y_{i}$ s are independent, moment-generating function of their sum is the product of their individual moment-generating functions, so

$$
m_{\sum Y_{i}}(t)=\prod_{i=1}^{9} m_{y}(t)=(1-\theta t)^{-9}
$$

and the moment-generating function of $2 / \theta$ times the sum of the $Y_{i} \mathrm{~S}$ is:

$$
m_{2 / \theta \sum}(t)=m_{\sum Y_{i}}(2 t / \theta)=(1-2 t)^{-9}
$$

which from the tables in the back of the text is the moment-generating function of a Chi-square variable with 18 degrees of freedom.
b) Use the pivotal quantity from part a) to derive a $95 \%$ confidence interval for $\theta$.

As before we start by finding $L$ and $U$ such that

$$
P\left(L \leq \chi_{18}^{2} \leq U\right)=0.95
$$

Either from the table in the text, or using
$L=\operatorname{CHIINV}(0.975,18)=8.23$ and $U=\operatorname{CHIINV}(0.025,18)=31.53$
we write

$$
P\left(8.23 \leq \frac{2 \sum_{i=1}^{9} Y_{i}}{\theta} \leq 31.53\right)=0.95
$$

which rearranges to

$$
P\left(\frac{2 \sum_{i=1}^{9} Y_{i}}{31.53} \leq \theta \leq \frac{2 \sum_{i=1}^{9} Y_{i}}{8.23}\right)=0.95
$$

5) Let $\left(Y_{1}, \ldots, Y_{5}\right)$ be a sample of size $n=5$ from a gamma distribution with $\alpha=2$ and $\beta$ unknown.
a) Use the method of moment-generating functions to show that

$$
\frac{2}{\beta} \sum_{i=1}^{5} Y_{i}
$$

is a pivotal quantity having a $\chi^{2}$ distribution with 20 degrees of freedom.

Because the $Y_{i}$ s are independent, moment-generating function of their sum is the product of their individual moment-generating functions, so

$$
m_{\sum Y_{i}}(t)=\prod_{i=1}^{5} m_{y}(t)=(1-\beta t)^{-2}
$$

and the moment-generating function of $2 / \beta$ times the sum of the $Y_{i} \mathrm{~S}$ is:

$$
m_{2 / \beta \sum}(t)=m_{\sum Y_{i}}(2 t / \beta)=(1-2 t)^{-10}
$$

which from the tables in the back of the text is the moment-generating function of a Chi-square variable with 20 degrees of freedom.
b) Use the pivotal quantity from part a) to derive a $95 \%$ confidence interval for $\theta$.
As before we start by finding $L$ and $U$ such that

$$
P\left(L \leq \chi_{20}^{2} \leq U\right)=0.95
$$

Either from the table in the text, or using
$L=\operatorname{CHIINV}(0.975,20)=9.59$ and $U=\operatorname{CHIINV}(0.025,20)=34.17$
we write

$$
P\left(9.59 \leq \frac{2 \sum_{i=1}^{5} Y_{i}}{\beta} \leq 34.17\right)=0.95
$$

which rearranges to

$$
P\left(\frac{2 \sum_{i=1}^{5} Y_{i}}{34.17} \leq \beta \leq \frac{2 \sum_{i=1}^{5} Y_{i}}{9.59}\right)=0.95
$$

