

**Name:**

**1)** Suppose  $(Y_1, \dots, Y_5)$  is a random sample of size 5 from a  $N(\mu, 1)$  distribution, and that

$$\hat{\mu}_1 = Y_1 + Y_3 + Y_5 - Y_2 - Y_4 \quad \text{and} \quad \hat{\mu}_2 = \bar{Y} = \frac{1}{5} \sum_{i=1}^5 Y_i$$

Show that  $\hat{\mu}_1$  and  $\hat{\mu}_2$  are unbiased and find the efficiency of  $\hat{\mu}_1$  relative to  $\hat{\mu}_2$ .

Note that both estimators are linear combinations of the sample values, with

$$t'_1 = (1, -1, 1, -1, 1) \quad \text{and} \quad t'_2 = \left( \frac{1}{5}, \frac{1}{5}, \frac{1}{5}, \frac{1}{5}, \frac{1}{5} \right)$$

Since

$$\mu' = (\mu, \mu, \mu, \mu, \mu) \quad \text{and} \quad V = I$$

we get

$$E(\hat{\mu}_1) = t'_1 \mu = 3\mu - 2\mu = \mu \quad \text{and} \quad E(\hat{\mu}_2) = t'_2 \mu = 5 \frac{\mu}{5} = \mu$$

so both are unbiased and

$$V(\hat{\mu}_1) = t'_1 V t_1 = t'_1 t_1 = 5 \quad \text{and} \quad V(\hat{\mu}_2) = t'_2 t_2 = 5 \frac{1}{5^2} = \frac{1}{5}$$

and the relative efficiency is

$$\text{eff}(\hat{\mu}_1, \hat{\mu}_2) = \frac{\hat{\mu}_1}{\hat{\mu}_2} = \frac{1}{25}$$

**2)** If  $Y$  has a binomial distribution with  $n$  trials and probability of success  $p$ , show that  $Y/n$  is a consistent estimator for  $p$ .

From Table 8.1, recall that

$$V(\hat{p}) = V(Y/n) = \frac{pq}{n} \rightarrow 0 \quad \text{as} \quad n \rightarrow \infty$$

So by Theorem 9.1 this estimator is consistent.

3) Suppose  $Y_1, Y_2, \dots, Y_n$  is a random sample and each  $Y_i$  has density function

$$f(y) = \begin{cases} \theta y^{\theta-1} & \text{if } 0 \leq y \leq 1 \\ 0 & \text{otherwise} \end{cases}$$

Show that  $\bar{Y}$  is a consistent estimator for  $\theta/(\theta + 1)$ .

Comparing the density function of  $Y$  to a beta distribution, we see that  $Y$  has a beta density with  $\alpha = \theta$  and  $\beta = 1$ .

From the results in the back cover of the text,

$$E(Y) = \frac{\alpha}{\alpha + \beta} = \frac{\theta}{\theta + 1} \quad \text{and} \quad V(Y) = \frac{\theta}{(\theta + 2)(\theta + 1)^2}$$

Using the results for expected values and variances of linear combinations, which are true regardless of the distribution of  $Y$ , we can say that:

$$E(\bar{Y}) = E(Y) = \frac{\theta}{\theta + 1} \quad \text{and} \quad V(\bar{Y}) = \frac{\theta}{n(\theta + 2)(\theta + 1)^2}$$

Since  $\bar{Y}$  is unbiased and its variance approaches zero as the sample size becomes large,  $\bar{Y}$  is a consistent estimator for  $E(Y) = \theta/(\theta + 1)$ .

4) Let  $Y_1, Y_2, \dots, Y_n$  be a random sample from a gamma distribution with parameters  $\alpha$  and  $\beta$ . Show that  $\bar{Y}$  converges in probability to a constant and find that constant.

From our results for linear combinations of random variables, we know that

$$E(\bar{Y}) = E(Y) = \alpha\beta$$

and

$$V(\bar{Y}) = \frac{1}{n}V(Y)$$

Since  $\bar{Y}$  is an unbiased estimate and its variance tends to zero,  $\bar{Y}$  is a consistent estimator for  $\alpha\beta$ , so  $\bar{Y}$  converges in probability to  $\alpha\beta$ .

5) Suppose  $Y_1, Y_2, \dots, Y_n$  is a random sample from the probability density function

$$f(y) = \begin{cases} \frac{2}{y^2} & \text{if } y \geq 2 \\ 0 & \text{otherwise} \end{cases}$$

Determine whether or not the law of large numbers applies to  $\bar{Y}$ .

Note that

$$E(Y^2) = \int_2^{\infty} y^2 \cdot \frac{2}{y^2} dy$$

Since the integral diverges,  $E(Y^2)$  is infinite so the variance of  $Y$  does not exist and the law of large numbers does not apply.